

Motivation: Chi-Square Distributions

Engineering Statistics

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PART I:

Motivation

χ_1^2 from $N(0, 1)$ cdf

Degrees of Freedom

χ^2 Distribution (Motivation)

Motivation:

- 1 Typical errors are distributed as $\text{Normal}(\mu, \sigma^2)$.
- 2 Standardized errors are distributed as $N(0, 1)$.
- 3 If $Z_1, \dots, Z_n \stackrel{iid}{\sim} N(0, 1)$, then their sum of squares $\sum_k Z_k^2$ is of interest...
...but how is this sum of squares of standard normal rv's distributed?

χ_1^2 Distribution from Standard Normal cdf

Theorem

(χ_1^2 Theorem – CHISQ1THM)

$$Z \sim N(0, 1) \implies Y := Z^2 \sim \chi_1^2 \text{ where } f_Y(y) = \frac{1}{\sqrt{2\pi}} y^{-1/2} e^{-y/2} = \frac{y^{1/2-1} e^{-y/2}}{2^{1/2} \cdot \Gamma(1/2)} \leftarrow \Gamma_{1/2, 2} \text{ pdf}$$

This distribution of Z^2 is called the **chi-square distribution with one degree of freedom**.

PROOF:

$$\begin{aligned} Z \sim N(0, 1) \\ \Omega_Z = (-\infty, \infty) \implies f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2} \implies \Phi(z) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^z e^{-t^2/2} dt \end{aligned}$$

Apply change of random variables (CRV): Let $Y := Z^2$

$$F_Y(y) = \mathbb{P}[Y \leq y] \stackrel{\text{CRV}}{=} \mathbb{P}[Z^2 \leq y] = \mathbb{P}[-\sqrt{y} \leq Z \leq \sqrt{y}] = \mathbb{P}[Z \leq \sqrt{y}] - \mathbb{P}[Z \leq -\sqrt{y}]$$

$$\stackrel{N(0,1)}{=} \Phi(\sqrt{y}) - \Phi(-\sqrt{y}) = \int_{-\infty}^{\sqrt{y}} \frac{1}{\sqrt{2\pi}} e^{-t^2/2} dt - \int_{-\infty}^{-\sqrt{y}} \frac{1}{\sqrt{2\pi}} e^{-t^2/2} dt = \int_{-\sqrt{y}}^{\sqrt{y}} \frac{1}{\sqrt{2\pi}} e^{-t^2/2} dt$$

$$\stackrel{\text{SYMI}}{=} 2 \cdot \int_0^{\sqrt{y}} \frac{1}{\sqrt{2\pi}} e^{-t^2/2} dt \stackrel{\text{CV}}{=} 2 \cdot \int_0^y \frac{1}{\sqrt{2\pi}} e^{-x/2} \cdot \frac{1}{2\sqrt{x}} dx = \int_0^y \frac{1}{\sqrt{2\pi}} x^{-1/2} e^{-x/2} dx$$

$$\text{(CV): Let } x := t^2 \implies dx = 2t dt \implies dt = \frac{1}{2\sqrt{x}} dx \implies \begin{cases} x(\sqrt{y}) &= & (\sqrt{y})^2 &= & y \\ x(0) &= & (0)^2 &= & 0 \end{cases}$$

χ_1^2 Distribution from Standard Normal cdf

Theorem

(χ_1^2 Theorem – CHISQ1THM)

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This distribution of Z^2 is called the **chi-square distribution with one degree of freedom**.

$$\begin{aligned} Z \sim N(0, 1) \\ \Omega_Z = (-\infty, \infty) \end{aligned} \implies f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2} \implies \Phi(z) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^z e^{-t^2/2} dt$$

Change of random variables (CRV): Let $Y := Z^2$

$$\implies \text{cdf of } Y \text{ is } F_Y(y) = \int_0^y \frac{1}{\sqrt{2\pi}} x^{-1/2} e^{-x/2} dx = \int_0^y \frac{1}{2^{1/2} \cdot \Gamma(1/2)} x^{1/2-1} e^{-x/2} dx$$

Limits of integration and integrand for cdf of $Y \implies \Omega_Y = (0, \infty)$

$$\implies \text{pdf of } Y \text{ is } f_Y(y) = \frac{1}{2^{1/2} \cdot \Gamma(1/2)} y^{1/2-1} e^{-y/2} \text{ which is the pdf of Gamma}(\alpha = 1/2, \beta = 2)$$

$$\therefore \begin{aligned} Z \sim N(0, 1) \\ \Omega_Z = (-\infty, \infty) \end{aligned} \implies \begin{aligned} Y := Z^2 \sim \chi_1^2 \\ \Omega_Y = (0, \infty) \end{aligned} \quad \square$$

R.V. Hogg, A.T. Craig, *Introduction to Mathematical Statistics*, 5th Ed, Prentice Hall, 1995. (§6.6)

R.J. Larsen, M.L. Marx, *An Intro to Mathematical Statistics...*, 2nd Ed, Prentice Hall, 1986. (§7.5)

A. Hald, *Statistical Theory with Engineering Applications*, Wiley, 1952. (§10.4)

Degrees of Freedom (General Motivation)

“Suppose you are asked to write 3 numbers with no restrictions upon them. You have complete freedom of choice in regard to all 3. There are 3 degrees of freedom.”[†]

“Now suppose you are asked to write 3 numbers with the restriction that their sum is to be some particular value, say 20. You cannot now choose all 3 freely, but as soon as 2 have been chosen the third is determined. Your choices are governed by the necessary relation $X_1 + X_2 + X_3 = 20$. In this situation there are only 2 degrees of freedom. The number of variables is 3, but the number of restrictions upon them is 1, and the number of ‘free’ variables, or independent choices, is $3 - 1 = 2$.”[†]

[†]H.M. Walker, J. Lev, *Statistical Inference*, Henry Holt and Company, 1953. (Ch 4)

Degrees of Freedom (General Motivation)

“Now suppose you are asked to write 5 numbers such that their sum is 30 and also such that the sum of the first two is 18. There are 5 variables but you do not have freedom of choice with respect to all 5. You cannot write 5 numbers arbitrarily and have them conform to the 2 restrictions that:

$$X_1 + X_2 = 18 \quad \text{and} \quad X_1 + X_2 + X_3 + X_4 + X_5 = 30$$

As soon as you select X_1 , then $X_2 = 18 - X_1$ and is completely determined. Since $X_3 + X_4 + X_5 = 30 - 18 = 12$, only two of the numbers, X_3 , X_4 , and X_5 , can be freely chosen. As one of the numbers X_1 and X_2 can be freely chosen there are 3 free choices. The number of degrees of freedom is $n = 5 - 2 = 3$.”[†]

“In every statistical problem in which degrees of freedom are involved it is necessary to determine the number of **free variables** by first noting the total number of variables and reducing that number by the number of **independent restrictions** upon them. In the preceding paragraph, for instance, one might think there are 3 restrictions, namely:

$$X_1 + X_2 = 18 \quad X_3 + X_4 + X_5 = 12 \quad X_1 + X_2 + X_3 + X_4 + X_5 = 30$$

However only two of these are independent, since any one of them can be deduced from the other two.”[†]

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χ^2 Distribution (Degrees of Freedom)

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...but how is this sum of squares of standard normal rv's distributed?

Since Z_1, \dots, Z_n are all independent,
there are no constraints imposed,
so they are n degrees of freedom.

$$\sum_k Z_k^2 \sim \chi_n^2$$

(chi-square with n degrees of freedom)

PART II:

Shearing Sum Mapping

 χ_2^2 from χ_1^2 & χ_1^2 pdf's χ_3^2 from χ_2^2 & χ_1^2 pdf's χ_4^2 from χ_3^2 & χ_1^2 pdf's χ_5^2 from χ_4^2 & χ_1^2 pdf's χ_6^2 from χ_5^2 & χ_1^2 pdf's χ_7^2 from χ_6^2 & χ_1^2 pdf's

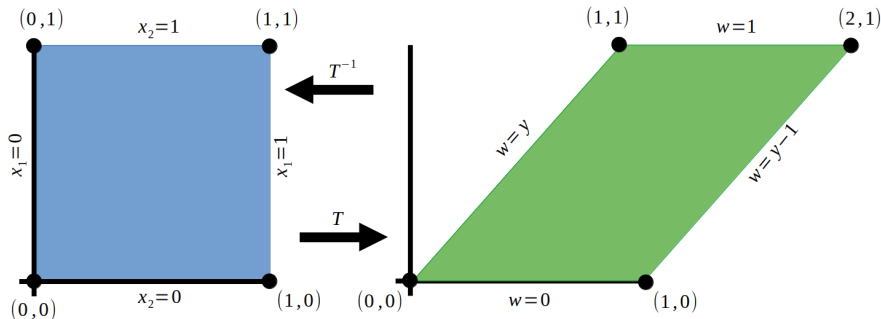
| | |
|-----------------|---------|
| χ_{2k}^2 | pattern |
| χ_{2k+1}^2 | pattern |

| | | | | | |
|-----------------|------|-----------------|---|------------|-------|
| χ_{2k}^2 | from | χ_{2k-1}^2 | & | χ_1^2 | pdf's |
| χ_{2k+1}^2 | from | χ_{2k}^2 | & | χ_1^2 | pdf's |

Adding Two RV's: Shearing Sum Mapping

$$T : \begin{cases} Y & := X_1 + X_2 \\ W & := X_2 \end{cases} \iff T^{-1} : \begin{cases} X_1 & := Y - W \\ X_2 & := W \end{cases}$$

| (x_1, x_2) | $(y, w) = T(x_1, x_2)$ | $(x_1, x_2) = T^{-1}(y, w)$ |
|--------------|---------------------------------|--------------------------------------|
| $(0, 0)$ | $T(0, 0) = (0 + 0, 0) = (0, 0)$ | $T^{-1}(0, 0) = (0 - 0, 0) = (0, 0)$ |
| $(1, 0)$ | $T(1, 0) = (1 + 0, 0) = (1, 0)$ | $T^{-1}(1, 0) = (1 - 0, 0) = (1, 0)$ |
| $(0, 1)$ | $T(0, 1) = (0 + 1, 1) = (1, 1)$ | $T^{-1}(1, 1) = (1 - 1, 1) = (0, 1)$ |
| $(1, 1)$ | $T(1, 1) = (1 + 1, 1) = (2, 1)$ | $T^{-1}(2, 1) = (2 - 1, 1) = (1, 1)$ |

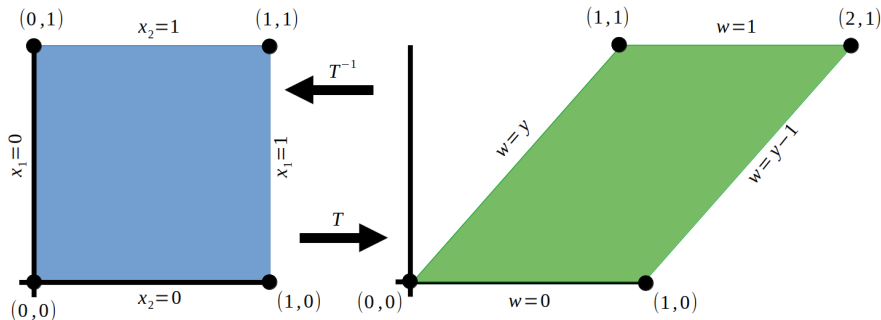


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Jacobian of inverse mapping $J_{T^{-1}} \equiv J = \begin{bmatrix} \frac{\partial X_1}{\partial Y} & \frac{\partial X_1}{\partial W} \\ \frac{\partial X_2}{\partial Y} & \frac{\partial X_2}{\partial W} \end{bmatrix} = \begin{bmatrix} 1 & -1 \\ 0 & 1 \end{bmatrix} \implies |J| = 1$

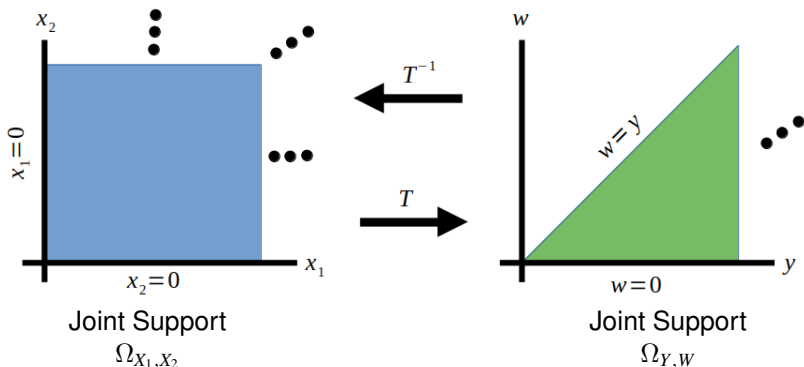
Jacobian equals one meaning the image of this shearing sum mapping of the unit square (with unit area) is the parallelogram also with the same unit area.



Adding Two RV's: Shearing Sum Mapping

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$$\Omega_{Y, W} \text{ is V-simple} \implies \Omega_1^V = \Omega_{Y, W}$$

$$\text{Top BC } \bar{\partial}\Omega_1^V = \{(y, w) : w = y\}$$

$$\text{Btm BC } \underline{\partial}\Omega_1^V = \{(y, w) : w = 0\}$$

χ_2^2 from χ_1^2 & χ_1^2 pdf's

Theorem

(χ_2^2 Theorem – CHISQ2THM)

$X_1, X_2 \stackrel{IID}{\sim} \chi_1^2 \implies Y := X_1 + X_2 \sim \chi_2^2$ where $f_Y(y) = \frac{1}{2}e^{-y/2} = \frac{y^{2/2-1}e^{-y/2}}{2^{2/2} \cdot \Gamma(2/2)} \leftarrow \Gamma_{1,2}$ pdf

This distribution of Y is called the **chi-square distribution with two degrees of freedom**.

PROOF:

pdf's $f_{X_1}(x_1) = \frac{x_1^{-1/2}e^{-x_1/2}}{\sqrt{2\pi}}$, $f_{X_2}(x_2) = \frac{x_2^{-1/2}e^{-x_2/2}}{\sqrt{2\pi}} \implies$ Supports $\Omega_{X_1} = \Omega_{X_2} = (0, \infty)$

Joint pdf $f_{X_1, X_2}(x_1, x_2) \stackrel{IND}{=} f_{X_1}(x_1) \cdot f_{X_2}(x_2) = \frac{x_1^{-1/2}e^{-x_1/2}}{\sqrt{2\pi}} \cdot \frac{x_2^{-1/2}e^{-x_2/2}}{\sqrt{2\pi}} = \frac{(x_1 x_2)^{-1/2}e^{-(x_1+x_2)/2}}{2\pi}$

Joint Support $\Omega_{X_1, X_2} \stackrel{IND}{=} \Omega_{X_1} \times \Omega_{X_2} = \{(x_1, x_2) : x_1 \in \Omega_{X_1}, x_2 \in \Omega_{X_2}\} = \{(x_1, x_2) : x_1 > 0, x_2 > 0\}$

Change of random variables (CRV):

$$\begin{aligned} Y := X_1 + X_2 \\ W := X_2 \end{aligned} \iff \begin{aligned} X_1 = Y - W \\ X_2 = W \end{aligned} \implies J = \begin{bmatrix} \frac{\partial X_1}{\partial Y} & \frac{\partial X_1}{\partial W} \\ \frac{\partial X_2}{\partial Y} & \frac{\partial X_2}{\partial W} \end{bmatrix} = \begin{bmatrix} 1 & -1 \\ 0 & 1 \end{bmatrix} \implies |J| = 1$$

\implies Joint pdf $f_{Y, W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2}w^{-1/2}e^{-y/2}}{2\pi}$

\implies Joint Support $\Omega_{Y, W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$

χ_2^2 from χ_1^2 & χ_1^2 pdf's

Theorem

(χ_2^2 Theorem – CHISQ2THM)

$X_1, X_2 \stackrel{iid}{\sim} \chi_1^2 \implies Y := X_1 + X_2 \sim \chi_2^2$ where $f_Y(y) = \frac{1}{2}e^{-y/2} = \frac{y^{2/2-1}e^{-y/2}}{2^{2/2} \cdot \Gamma(2/2)} \leftarrow \Gamma_{1,2}$ pdf

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PROOF: \implies Joint pdf $f_{Y,W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2} w^{-1/2} e^{-y/2}}{2\pi}$

\implies Joint Support $\Omega_{Y,W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$

\implies Marginal pdf $f_Y(y) = \int_{\underline{\partial}\Omega_Y^Y}^{\overline{\partial}\Omega_Y^Y} f_{Y,W}(y, w) dw = \int_0^y \frac{(y-w)^{-1/2} w^{-1/2} e^{-y/2}}{2\pi} dw$

$$f_Y(y) = \frac{e^{-y/2}}{2\pi} \int_0^y \frac{dw}{w^{1/2}(y-w)^{1/2}} \quad (CV) \quad w := y \sin^2 \theta \iff \begin{cases} y-w & = & y \cos^2 \theta \\ dw & = & 2y \sin \theta \cos \theta d\theta \\ w=y & \iff & \theta = \pi/2 \\ w=0 & \iff & \theta = 0 \end{cases}$$

$$f_Y(y) \stackrel{CV}{=} \frac{e^{-y/2}}{2\pi} \cdot \int_0^{\pi/2} \frac{2y \sin \theta \cos \theta d\theta}{y^{1/2} \sin \theta \cdot y^{1/2} \cos \theta} = \frac{e^{-y/2}}{\pi} \cdot \int_0^{\pi/2} d\theta = \frac{e^{-y/2}}{\pi} \cdot [\theta]_{\theta=0}^{\theta=\pi/2}$$

$$f_Y(y) \stackrel{FTC}{=} \frac{e^{-y/2}}{\pi} \cdot \left[\frac{\pi}{2} - 0 \right] = \frac{1}{2} e^{-y/2} \quad \therefore f_Y(y) = \frac{1}{2} e^{-y/2} = \frac{y^{2/2-1} e^{-y/2}}{2^{2/2} \cdot \Gamma(2/2)} \quad \square$$

H.W. Alexander, *Elements of Mathematical Statistics*, Wiley, 1961. (Ch5, §37)

χ_3^2 from χ_2^2 & χ_1^2 pdf's

Theorem

(χ_3^2 Theorem – CHISQ3THM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_2^2 \implies Y := X_1 + X_2 \sim \chi_3^2$ where $f_Y(y) = \frac{y^{1/2} e^{-y/2}}{\sqrt{2\pi}} = \frac{y^{3/2-1} e^{-y/2}}{2^{3/2} \cdot \Gamma(3/2)} \leftarrow \Gamma_{3/2,2}$ pdf

This distribution of Y is called the **chi-square distribution with three degrees of freedom**.

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Joint pdf $f_{X_1, X_2}(x_1, x_2) \stackrel{IND}{=} f_{X_1}(x_1) \cdot f_{X_2}(x_2) = \frac{x_1^{-1/2} e^{-x_1/2}}{\sqrt{2\pi}} \cdot \frac{e^{-x_2/2}}{2} = \frac{x_1^{-1/2} e^{-(x_1+x_2)/2}}{2\sqrt{2\pi}}$

Joint Support $\Omega_{X_1, X_2} \stackrel{IND}{=} \Omega_{X_1} \times \Omega_{X_2} = \{(x_1, x_2) : x_1 \in \Omega_{X_1}, x_2 \in \Omega_{X_2}\} = \{(x_1, x_2) : x_1 > 0, x_2 > 0\}$

Change of random variables (CRV):

$$\begin{aligned} Y := X_1 + X_2 \\ W := X_2 \end{aligned} \iff \begin{aligned} X_1 = Y - W \\ X_2 = W \end{aligned} \implies J = \begin{bmatrix} \frac{\partial X_1}{\partial Y} & \frac{\partial X_1}{\partial W} \\ \frac{\partial X_2}{\partial Y} & \frac{\partial X_2}{\partial W} \end{bmatrix} = \begin{bmatrix} 1 & -1 \\ 0 & 1 \end{bmatrix} \implies |J| = 1$$

\implies Joint pdf $f_{Y,W}(y,w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2} e^{-y/2}}{2\sqrt{2\pi}}$

\implies Joint Support $\Omega_{Y,W} = \{(y,w) : 0 < y < \infty, 0 < w < y\}$

χ_3^2 from χ_2^2 & χ_1^2 pdf's

Theorem

(χ_3^2 Theorem – CHISQ3THM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_2^2 \implies Y := X_1 + X_2 \sim \chi_3^2$ where $f_Y(y) = \frac{y^{1/2}e^{-y/2}}{\sqrt{2\pi}} = \frac{y^{3/2-1}e^{-y/2}}{2^{3/2} \cdot \Gamma(3/2)} \leftarrow \Gamma_{3/2,2}$ pdf

This distribution of Y is called the **chi-square distribution with three degrees of freedom**.

PROOF:

$$\implies \text{Joint pdf } f_{Y,W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2}e^{-y/2}}{2\sqrt{2\pi}}$$

$$\implies \text{Joint Support } \Omega_{Y,W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$$

$$\implies \text{Marginal pdf } f_Y(y) = \int_{\underline{\partial}\Omega_1^Y}^{\bar{\partial}\Omega_1^Y} f_{Y,W}(y, w) dw = \int_0^y \frac{(y-w)^{-1/2}e^{-y/2}}{2\sqrt{2\pi}} dw$$

$$f_Y(y) = \frac{e^{-y/2}}{2\sqrt{2\pi}} \int_0^y \frac{dw}{(y-w)^{1/2}} \quad (CV) \quad w := y \sin^2 \theta \iff \begin{cases} y-w & = & y \cos^2 \theta \\ dw & = & 2y \sin \theta \cos \theta d\theta \\ w=y & \iff & \theta = \pi/2 \\ w=0 & \iff & \theta = 0 \end{cases}$$

$$f_Y(y) \stackrel{CV}{=} \frac{e^{-y/2}}{2\sqrt{2\pi}} \cdot \int_0^{\pi/2} \frac{2y \sin \theta \cos \theta d\theta}{y^{1/2} \cos \theta} = \frac{y^{1/2}e^{-y/2}}{\sqrt{2\pi}} \cdot \int_0^{\pi/2} \sin \theta d\theta = \frac{y^{1/2}e^{-y/2}}{\sqrt{2\pi}} \cdot \left[-\cos \theta \right]_{\theta=0}^{\theta=\pi/2}$$

$$f_Y(y) \stackrel{FTC}{=} \frac{y^{1/2}e^{-y/2}}{\sqrt{2\pi}} \cdot [0 - (-1)] = \frac{1}{\sqrt{2\pi}} y^{1/2} e^{-y/2} \quad \therefore f_Y(y) = \frac{1}{\sqrt{2\pi}} y^{1/2} e^{-y/2} = \frac{y^{3/2-1} e^{-y/2}}{2^{3/2} \cdot \Gamma(3/2)} \quad \square$$

χ_4^2 from χ_3^2 & χ_1^2 pdf's

Theorem

(χ_4^2 Theorem – CHISQ4THM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_3^2 \implies Y := X_1 + X_2 \sim \chi_4^2$ where $f_Y(y) = \frac{1}{4}ye^{-y/2} = \frac{y^{4/2-1}e^{-y/2}}{2^{4/2} \cdot \Gamma(4/2)} \leftarrow \Gamma_{2,2}$ pdf

This distribution of Y is called the **chi-square distribution with four degrees of freedom**.

PROOF:

pdf's $f_{X_1}(x_1) = \frac{x_1^{-1/2}e^{-x_1/2}}{\sqrt{2\pi}}$, $f_{X_2}(x_2) = \frac{x_2^{1/2}e^{-x_2/2}}{\sqrt{2\pi}} \implies$ Supports $\Omega_{X_1} = \Omega_{X_2} = (0, \infty)$

Joint pdf $f_{X_1, X_2}(x_1, x_2) \stackrel{IND}{=} f_{X_1}(x_1) \cdot f_{X_2}(x_2) = \frac{x_1^{-1/2}e^{-x_1/2}}{\sqrt{2\pi}} \cdot \frac{x_2^{1/2}e^{-x_2/2}}{\sqrt{2\pi}} = \frac{x_1^{-1/2}x_2^{1/2}e^{-(x_1+x_2)/2}}{2\pi}$

Joint Support $\Omega_{X_1, X_2} \stackrel{IND}{=} \Omega_{X_1} \times \Omega_{X_2} = \{(x_1, x_2) : x_1 \in \Omega_{X_1}, x_2 \in \Omega_{X_2}\} = \{(x_1, x_2) : x_1 > 0, x_2 > 0\}$

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\implies Joint pdf $f_{Y, W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2}w^{1/2}e^{-y/2}}{2\pi}$

\implies Joint Support $\Omega_{Y, W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$

χ_4^2 from χ_3^2 & χ_1^2 pdf's

Theorem

(χ_4^2 Theorem – CHISQ4THM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_3^2 \implies Y := X_1 + X_2 \sim \chi_4^2$ where $f_Y(y) = \frac{1}{4}ye^{-y/2} = \frac{y^{4/2-1}e^{-y/2}}{2^{4/2} \cdot \Gamma(4/2)} \leftarrow \Gamma_{2,2}$ pdf

This distribution of Y is called the **chi-square distribution with four degrees of freedom**.

PROOF:

$$\implies \text{Joint pdf } f_{Y,W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2} w^{1/2} e^{-y/2}}{2\pi}$$

$$\implies \text{Joint Support } \Omega_{Y,W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$$

$$\implies \text{Marginal pdf } f_Y(y) = \int_{\underline{\partial\Omega}_Y^y}^{\overline{\partial\Omega}_Y^y} f_{Y,W}(y, w) dw = \int_0^y \frac{(y-w)^{-1/2} w^{1/2} e^{-y/2}}{2\pi} dw$$

$$f_Y(y) = \frac{e^{-y/2}}{2\pi} \int_0^y \frac{w^{1/2} dw}{(y-w)^{1/2}} \quad (CV) \quad w := y \sin^2 \theta \iff \begin{cases} y-w & = & y \cos^2 \theta \\ dw & = & 2y \sin \theta \cos \theta d\theta \\ w=y & \iff & \theta = \pi/2 \\ w=0 & \iff & \theta = 0 \end{cases}$$

$$f_Y(y) \stackrel{CV}{=} \frac{e^{-y/2}}{2\pi} \cdot \int_0^{\pi/2} \frac{(y^{1/2} \sin \theta)(2y \sin \theta \cos \theta d\theta)}{y^{1/2} \cos \theta}$$

$$f_Y(y) = \frac{ye^{-y/2}}{\pi} \cdot \int_0^{\pi/2} \sin^2 \theta d\theta$$

χ_4^2 from χ_3^2 & χ_1^2 pdf's

Theorem

(χ_4^2 Theorem – CHISQ4THM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_3^2 \implies Y := X_1 + X_2 \sim \chi_4^2$ where $f_Y(y) = \frac{1}{4}ye^{-y/2} = \frac{y^{4/2-1}e^{-y/2}}{2^{4/2} \cdot \Gamma(4/2)} \leftarrow \Gamma_{2,2}$ pdf

This distribution of Y is called the **chi-square distribution with four degrees of freedom**.

PROOF:

$$\implies \text{Joint pdf } f_{Y,W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2}w^{1/2}e^{-y/2}}{2\pi}$$

$$\implies \text{Joint Support } \Omega_{Y,W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$$

$$\implies \text{Marginal pdf } f_Y(y) = \int_{\underline{\partial}\Omega_Y^y}^{\overline{\partial}\Omega_Y^y} f_{Y,W}(y, w) dw = \int_0^y \frac{(y-w)^{-1/2}w^{1/2}e^{-y/2}}{2\pi} dw$$

$$f_Y(y) = \frac{ye^{-y/2}}{\pi} \cdot \int_0^{\pi/2} \sin^2 \theta d\theta \stackrel{TRIG}{=} \frac{ye^{-y/2}}{\pi} \cdot \int_0^{\pi/2} \frac{1}{2} - \frac{\cos 2\theta}{2} d\theta = \frac{ye^{-y/2}}{\pi} \cdot \left[\frac{\theta}{2} - \frac{\sin 2\theta}{4} \right]_{\theta=0}^{\theta=\pi/2}$$

$$f_Y(y) \stackrel{FTC}{=} \frac{ye^{-y/2}}{\pi} \cdot \left[\left(\frac{\pi}{4} - 0 \right) - \left(\frac{0}{2} - 0 \right) \right] = \frac{1}{4}ye^{-y/2} \quad \therefore f_Y(y) = \frac{1}{4}ye^{-y/2} = \frac{y^{4/2-1}e^{-y/2}}{2^{4/2} \cdot \Gamma(4/2)} \quad \square$$

χ_5^2 from χ_4^2 & χ_1^2 pdf's

Theorem

(χ_5^2 Theorem – CHISQ5THM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_4^2 \implies Y := X_1 + X_2 \sim \chi_5^2$ where $f_Y(y) = \frac{y^{3/2} e^{-y/2}}{3\sqrt{2\pi}} = \frac{y^{5/2-1} e^{-y/2}}{2^{5/2} \cdot \Gamma(5/2)} \leftarrow \Gamma_{5/2,2}$ pdf

This distribution of Y is called the **chi-square distribution with five degrees of freedom**.

PROOF:

pdf's $f_{X_1}(x_1) = \frac{x_1^{-1/2} e^{-x_1/2}}{\sqrt{2\pi}}$, $f_{X_2}(x_2) = \frac{1}{4} x_2 e^{-x_2/2} \implies$ Supports $\Omega_{X_1} = \Omega_{X_2} = (0, \infty)$

Joint pdf $f_{X_1, X_2}(x_1, x_2) \stackrel{IND}{=} f_{X_1}(x_1) \cdot f_{X_2}(x_2) = \frac{x_1^{-1/2} e^{-x_1/2}}{\sqrt{2\pi}} \cdot \frac{1}{4} x_2 e^{-x_2/2} = \frac{x_1^{-1/2} x_2 e^{-(x_1+x_2)/2}}{4\sqrt{2\pi}}$

Joint Support $\Omega_{X_1, X_2} \stackrel{IND}{=} \Omega_{X_1} \times \Omega_{X_2} = \{(x_1, x_2) : x_1 \in \Omega_{X_1}, x_2 \in \Omega_{X_2}\} = \{(x_1, x_2) : x_1 > 0, x_2 > 0\}$

Change of random variables (CRV):

$$\begin{aligned} Y := X_1 + X_2 \\ W := X_2 \end{aligned} \iff \begin{aligned} X_1 = Y - W \\ X_2 = W \end{aligned} \implies J = \begin{bmatrix} \frac{\partial X_1}{\partial Y} & \frac{\partial X_1}{\partial W} \\ \frac{\partial X_2}{\partial Y} & \frac{\partial X_2}{\partial W} \end{bmatrix} = \begin{bmatrix} 1 & -1 \\ 0 & 1 \end{bmatrix} \implies |J| = 1$$

\implies Joint pdf $f_{Y,W}(y,w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2} w e^{-y/2}}{4\sqrt{2\pi}}$

\implies Joint Support $\Omega_{Y,W} = \{(y,w) : 0 < y < \infty, 0 < w < y\}$

χ_5^2 from χ_4^2 & χ_1^2 pdf's

Theorem

(χ_5^2 Theorem – CHISQ5THM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_4^2 \implies Y := X_1 + X_2 \sim \chi_5^2$ where $f_Y(y) = \frac{y^{3/2}e^{-y/2}}{3\sqrt{2\pi}} = \frac{y^{5/2-1}e^{-y/2}}{2^{5/2} \cdot \Gamma(5/2)} \leftarrow \Gamma_{5/2,2}$ pdf

This distribution of Y is called the **chi-square distribution with five degrees of freedom**.

PROOF:

$$\implies \text{Joint pdf } f_{Y,W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2} w e^{-y/2}}{4\sqrt{2\pi}}$$

$$\implies \text{Joint Support } \Omega_{Y,W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$$

$$\implies \text{Marginal pdf } f_Y(y) = \int_{\underline{\partial}\Omega_1^Y}^{\bar{\partial}\Omega_1^Y} f_{Y,W}(y, w) dw = \int_0^y \frac{(y-w)^{-1/2} w e^{-y/2}}{4\sqrt{2\pi}} dw$$

$$f_Y(y) = \frac{e^{-y/2}}{4\sqrt{2\pi}} \int_0^y \frac{w dw}{(y-w)^{1/2}} \quad (CV) \quad w := y \sin^2 \theta \iff \begin{cases} y-w & = & y \cos^2 \theta \\ dw & = & 2y \sin \theta \cos \theta d\theta \\ w=y & \iff & \theta = \pi/2 \\ w=0 & \iff & \theta = 0 \end{cases}$$

$$f_Y(y) \stackrel{CV}{=} \frac{e^{-y/2}}{4\sqrt{2\pi}} \cdot \int_0^{\pi/2} \frac{(y \sin^2 \theta)(2y \sin \theta \cos \theta d\theta)}{y^{1/2} \cos \theta}$$

$$f_Y(y) = \frac{y^{3/2} e^{-y/2}}{2\sqrt{2\pi}} \cdot \int_0^{\pi/2} \sin^3 \theta d\theta$$

χ_5^2 from χ_4^2 & χ_1^2 pdf's

Theorem

(χ_5^2 Theorem – CHISQ5THM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_4^2 \implies Y := X_1 + X_2 \sim \chi_5^2$ where $f_Y(y) = \frac{y^{3/2} e^{-y/2}}{3\sqrt{2\pi}} = \frac{y^{5/2-1} e^{-y/2}}{2^{5/2} \cdot \Gamma(5/2)} \leftarrow \Gamma_{5/2,2}$ pdf

This distribution of Y is called the **chi-square distribution with five degrees of freedom**.

PROOF:

$$\implies \text{Joint pdf } f_{Y,W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2} w e^{-y/2}}{4\sqrt{2\pi}}$$

$$\implies \text{Joint Support } \Omega_{Y,W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$$

$$\implies \text{Marginal pdf } f_Y(y) = \int_{\underline{\partial}\Omega_Y^y}^{\overline{\partial}\Omega_Y^y} f_{Y,W}(y, w) dw = \int_0^y \frac{(y-w)^{-1/2} w e^{-y/2}}{4\sqrt{2\pi}} dw$$

$$f_Y(y) = \frac{y^{3/2} e^{-y/2}}{2\sqrt{2\pi}} \cdot \int_0^{\pi/2} \sin^3 \theta d\theta \stackrel{TRIG}{=} \frac{y^{3/2} e^{-y/2}}{2\sqrt{2\pi}} \cdot \int_0^{\pi/2} (1 - \cos^2 \theta) \sin \theta d\theta \quad (CV) \quad u := \cos \theta$$

$$f_Y(y) \stackrel{CV}{=} \frac{y^{3/2} e^{-y/2}}{2\sqrt{2\pi}} \cdot \int_0^1 (1 - u^2) du = \frac{y^{3/2} e^{-y/2}}{2\sqrt{2\pi}} \cdot \left[u - \frac{u^3}{3} \right]_{u=0}^{u=1}$$

$$\stackrel{FTC}{=} \frac{y^{3/2} e^{-y/2}}{2\sqrt{2\pi}} \cdot \left[\left(1 - \frac{1^3}{3}\right) - \left(0 - \frac{0^3}{3}\right) \right] = \frac{y^{3/2} e^{-y/2}}{2\sqrt{2\pi}} \cdot \frac{2}{3} = \frac{y^{3/2} e^{-y/2}}{3\sqrt{2\pi}} = \frac{y^{5/2-1} e^{-y/2}}{2^{5/2} \cdot \Gamma(5/2)} \quad \square$$

χ_6^2 from χ_5^2 & χ_1^2 pdf's

Theorem

(χ_6^2 Theorem – CHISQ6THM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_5^2 \implies Y := X_1 + X_2 \sim \chi_6^2$ where $f_Y(y) = \frac{1}{16} y^2 e^{-y/2} = \frac{y^{6/2-1} e^{-y/2}}{2^{6/2} \cdot \Gamma(6/2)} \leftarrow \Gamma_{3,2}$ pdf

This distribution of Y is called the **chi-square distribution with six degrees of freedom**.

PROOF:

pdf's $f_{X_1}(x_1) = \frac{x_1^{-1/2} e^{-x_1/2}}{\sqrt{2\pi}}$, $f_{X_2}(x_2) = \frac{x_2^{3/2} e^{-x_2/2}}{3\sqrt{2\pi}} \implies$ Supports $\Omega_{X_1} = \Omega_{X_2} = (0, \infty)$

Joint pdf $f_{X_1, X_2}(x_1, x_2) \stackrel{IND}{=} f_{X_1}(x_1) \cdot f_{X_2}(x_2) = \frac{x_1^{-1/2} e^{-x_1/2}}{\sqrt{2\pi}} \cdot \frac{x_2^{3/2} e^{-x_2/2}}{3\sqrt{2\pi}} = \frac{x_1^{-1/2} x_2^{3/2} e^{-(x_1+x_2)/2}}{6\pi}$

Joint Support $\Omega_{X_1, X_2} \stackrel{IND}{=} \Omega_{X_1} \times \Omega_{X_2} = \{(x_1, x_2) : x_1 \in \Omega_{X_1}, x_2 \in \Omega_{X_2}\} = \{(x_1, x_2) : x_1 > 0, x_2 > 0\}$

Change of random variables (CRV):

$$\begin{aligned} Y := X_1 + X_2 & \iff X_1 = Y - W \\ W := X_2 & \iff X_2 = W \end{aligned} \implies J = \begin{bmatrix} \frac{\partial X_1}{\partial Y} & \frac{\partial X_1}{\partial W} \\ \frac{\partial X_2}{\partial Y} & \frac{\partial X_2}{\partial W} \end{bmatrix} = \begin{bmatrix} 1 & -1 \\ 0 & 1 \end{bmatrix} \implies |J| = 1$$

\implies Joint pdf $f_{Y, W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2} w^{3/2} e^{-y/2}}{6\pi}$

\implies Joint Support $\Omega_{Y, W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$

χ_6^2 from χ_5^2 & χ_1^2 pdf's

Theorem

(χ_6^2 Theorem – CHISQ6THM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_5^2 \implies Y := X_1 + X_2 \sim \chi_6^2$ where $f_Y(y) = \frac{1}{16}y^2e^{-y/2} = \frac{y^{6/2-1}e^{-y/2}}{2^{6/2} \cdot \Gamma(6/2)} \leftarrow \Gamma_{3,2}$ pdf

This distribution of Y is called the **chi-square distribution with six degrees of freedom**.

PROOF:

$$\implies \text{Joint pdf } f_{Y,W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2}w^{3/2}e^{-y/2}}{6\pi}$$

$$\implies \text{Joint Support } \Omega_{Y,W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$$

$$\implies \text{Marginal pdf } f_Y(y) = \int_{\underline{\partial\Omega}_1^Y}^{\overline{\partial\Omega}_1^Y} f_{Y,W}(y, w) dw = \int_0^y \frac{(y-w)^{-1/2}w^{3/2}e^{-y/2}}{6\pi} dw$$

$$f_Y(y) = \frac{e^{-y/2}}{6\pi} \int_0^y \frac{w^{3/2} dw}{(y-w)^{1/2}} \quad (CV) \quad w := y \sin^2 \theta \iff \begin{cases} y-w & = & y \cos^2 \theta \\ dw & = & 2y \sin \theta \cos \theta d\theta \\ w=y & \iff & \theta = \pi/2 \\ w=0 & \iff & \theta = 0 \end{cases}$$

$$f_Y(y) \stackrel{CV}{=} \frac{e^{-y/2}}{6\pi} \cdot \int_0^{\pi/2} \frac{(y^{3/2} \sin^3 \theta)(2y \sin \theta \cos \theta d\theta)}{y^{1/2} \cos \theta}$$

$$f_Y(y) = \frac{y^2 e^{-y/2}}{3\pi} \cdot \int_0^{\pi/2} \sin^4 \theta d\theta \stackrel{TRIG}{=} \frac{y^2 e^{-y/2}}{3\pi} \cdot \int_0^{\pi/2} \left(\frac{1}{2} - \frac{1}{2} \cos 2\theta \right)^2 d\theta$$

χ_6^2 from χ_5^2 & χ_1^2 pdf's

Theorem

(χ_6^2 Theorem – CHISQ6THM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_5^2 \implies Y := X_1 + X_2 \sim \chi_6^2$ where $f_Y(y) = \frac{1}{16} y^2 e^{-y/2} = \frac{y^{6/2-1} e^{-y/2}}{2^{6/2} \cdot \Gamma(6/2)} \leftarrow \Gamma_{3,2}$ pdf

This distribution of Y is called the **chi-square distribution with six degrees of freedom**.

PROOF:

$$\implies \text{Joint pdf } f_{Y,W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2} w^{3/2} e^{-y/2}}{6\pi}$$

$$\implies \text{Joint Support } \Omega_{Y,W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$$

$$\implies \text{Marginal pdf } f_Y(y) = \int_{\underline{\partial\Omega}_Y^y}^{\overline{\partial\Omega}_Y^y} f_{Y,W}(y, w) dw = \int_0^y \frac{(y-w)^{-1/2} w^{3/2} e^{-y/2}}{6\pi} dw$$

$$f_Y(y) = \frac{y^2 e^{-y/2}}{3\pi} \cdot \int_0^{\pi/2} \left(\frac{1}{2} - \frac{1}{2} \cos 2\theta \right)^2 d\theta = \frac{y^2 e^{-y/2}}{3\pi} \cdot \int_0^{\pi/2} \left(\frac{1}{4} - \frac{1}{2} \cos 2\theta + \frac{1}{4} \cos^2 2\theta \right) d\theta$$

$$f_Y(y) \stackrel{TRIG}{=} \frac{y^2 e^{-y/2}}{3\pi} \cdot \int_0^{\pi/2} \left[\frac{1}{4} - \frac{1}{2} \cos 2\theta + \frac{1}{4} \left(\frac{1}{2} + \frac{1}{2} \cos 4\theta \right) \right] d\theta$$

$$f_Y(y) = \frac{y^2 e^{-y/2}}{3\pi} \cdot \int_0^{\pi/2} \left(\frac{3}{8} - \frac{1}{2} \cos 2\theta + \frac{1}{8} \cos 4\theta \right) d\theta$$

χ_6^2 from χ_5^2 & χ_1^2 pdf's

Theorem

(χ_6^2 Theorem – CHISQ6THM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_5^2 \implies Y := X_1 + X_2 \sim \chi_6^2$ where $f_Y(y) = \frac{1}{16}y^2e^{-y/2} = \frac{y^{6/2-1}e^{-y/2}}{2^{6/2} \cdot \Gamma(6/2)} \leftarrow \Gamma_{3,2}$ pdf

This distribution of Y is called the **chi-square distribution with six degrees of freedom**.

PROOF:

$$\implies \text{Joint pdf } f_{Y,W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2} w^{3/2} e^{-y/2}}{6\pi}$$

$$\implies \text{Joint Support } \Omega_{Y,W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$$

$$\implies \text{Marginal pdf } f_Y(y) = \int_{\partial\Omega_Y^y} f_{Y,W}(y, w) dw = \int_0^y \frac{(y-w)^{-1/2} w^{3/2} e^{-y/2}}{6\pi} dw$$

$$f_Y(y) = \frac{y^2 e^{-y/2}}{3\pi} \cdot \int_0^{\pi/2} \left(\frac{3}{8} - \frac{1}{2} \cos 2\theta + \frac{1}{8} \cos 4\theta \right) d\theta$$

$$f_Y(y) = \frac{y^2 e^{-y/2}}{3\pi} \cdot \left[\frac{3}{8}\theta - \frac{1}{4} \sin 2\theta + \frac{1}{32} \sin 4\theta \right]_{\theta=0}^{\theta=\pi/2}$$

$$f_Y(y) \stackrel{FTC}{=} \frac{y^2 e^{-y/2}}{3\pi} \cdot \frac{3\pi}{16} = \frac{1}{16} y^2 e^{-y/2} = \frac{y^{6/2-1} e^{-y/2}}{2^{6/2} \cdot \Gamma(6/2)} \quad \square$$

χ_7^2 from χ_6^2 & χ_1^2 pdf's

Theorem

(χ_7^2 Theorem – CHISQ7THM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_6^2 \implies Y := X_1 + X_2 \sim \chi_7^2$ where $f_Y(y) = \frac{y^{5/2} e^{-y/2}}{15\sqrt{2\pi}} = \frac{y^{7/2-1} e^{-y/2}}{2^{7/2} \cdot \Gamma(7/2)} \leftarrow \Gamma_{7/2,2}$ pdf

This distribution of Y is called the **chi-square distribution with seven degrees of freedom**.

PROOF:

pdf's $f_{X_1}(x_1) = \frac{x_1^{-1/2} e^{-x_1/2}}{\sqrt{2\pi}}$, $f_{X_2}(x_2) = \frac{1}{16} x_2^2 e^{-x_2/2} \implies$ Supports $\Omega_{X_1} = \Omega_{X_2} = (0, \infty)$

Joint pdf $f_{X_1, X_2}(x_1, x_2) \stackrel{IND}{=} f_{X_1}(x_1) \cdot f_{X_2}(x_2) = \frac{x_1^{-1/2} e^{-x_1/2}}{\sqrt{2\pi}} \cdot \frac{1}{16} x_2^2 e^{-x_2/2} = \frac{x_1^{-1/2} x_2^2 e^{-(x_1+x_2)/2}}{16\sqrt{2\pi}}$

Joint Support $\Omega_{X_1, X_2} \stackrel{IND}{=} \Omega_{X_1} \times \Omega_{X_2} = \{(x_1, x_2) : x_1 \in \Omega_{X_1}, x_2 \in \Omega_{X_2}\} = \{(x_1, x_2) : x_1 > 0, x_2 > 0\}$

Change of random variables (CRV):

$$\begin{aligned} Y := X_1 + X_2 & \iff X_1 = Y - W \\ W := X_2 & \iff X_2 = W \end{aligned} \implies J = \begin{bmatrix} \frac{\partial X_1}{\partial Y} & \frac{\partial X_1}{\partial W} \\ \frac{\partial X_2}{\partial Y} & \frac{\partial X_2}{\partial W} \end{bmatrix} = \begin{bmatrix} 1 & -1 \\ 0 & 1 \end{bmatrix} \implies |J| = 1$$

\implies Joint pdf $f_{Y,W}(y,w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2} w^2 e^{-y/2}}{16\sqrt{2\pi}}$

\implies Joint Support $\Omega_{Y,W} = \{(y,w) : 0 < y < \infty, 0 < w < y\}$

χ_7^2 from χ_6^2 & χ_1^2 pdf's

Theorem

(χ_7^2 Theorem – CHISQ7THM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_6^2 \implies Y := X_1 + X_2 \sim \chi_7^2$ where $f_Y(y) = \frac{y^{5/2} e^{-y/2}}{15\sqrt{2\pi}} = \frac{y^{7/2-1} e^{-y/2}}{2^{7/2} \cdot \Gamma(7/2)} \leftarrow \Gamma_{7/2,2}$ pdf

This distribution of Y is called the **chi-square distribution with seven degrees of freedom**.

PROOF:

$$\implies \text{Joint pdf } f_{Y,W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2} w^2 e^{-y/2}}{16\sqrt{2\pi}}$$

$$\implies \text{Joint Support } \Omega_{Y,W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$$

$$\implies \text{Marginal pdf } f_Y(y) = \int_{\underline{\partial}\Omega_1^Y}^{\bar{\partial}\Omega_1^Y} f_{Y,W}(y, w) dw = \int_0^y \frac{(y-w)^{-1/2} w^2 e^{-y/2}}{16\sqrt{2\pi}} dw$$

$$f_Y(y) = \frac{e^{-y/2}}{16\sqrt{2\pi}} \int_0^y \frac{w^2 dw}{(y-w)^{1/2}} \quad (CV) \quad w := y \sin^2 \theta \iff \begin{cases} y-w = & y \cos^2 \theta \\ dw = & 2y \sin \theta \cos \theta d\theta \\ w=y \iff & \theta = \pi/2 \\ w=0 \iff & \theta = 0 \end{cases}$$

$$f_Y(y) \stackrel{CV}{=} \frac{e^{-y/2}}{16\sqrt{2\pi}} \cdot \int_0^{\pi/2} \frac{(y^2 \sin^4 \theta)(2y \sin \theta \cos \theta d\theta)}{y^{1/2} \cos \theta}$$

$$f_Y(y) = \frac{y^{5/2} e^{-y/2}}{8\sqrt{2\pi}} \cdot \int_0^{\pi/2} \sin^5 \theta d\theta \stackrel{TRIG}{=} \frac{y^{5/2} e^{-y/2}}{8\sqrt{2\pi}} \cdot \int_0^{\pi/2} \left(\frac{1}{2} - \frac{1}{2} \cos 2\theta \right)^2 \sin \theta d\theta$$

χ_7^2 from χ_6^2 & χ_1^2 pdf's

Theorem

(χ_7^2 Theorem – CHISQ7THM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_6^2 \implies Y := X_1 + X_2 \sim \chi_7^2$ where $f_Y(y) = \frac{y^{5/2} e^{-y/2}}{15\sqrt{2\pi}} = \frac{y^{7/2-1} e^{-y/2}}{2^{7/2} \cdot \Gamma(7/2)} \leftarrow \Gamma_{7/2,2}$ pdf

This distribution of Y is called the **chi-square distribution with seven degrees of freedom**.

PROOF:

$$\implies \text{Joint pdf } f_{Y,W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2} w^2 e^{-y/2}}{16\sqrt{2\pi}}$$

$$\implies \text{Joint Support } \Omega_{Y,W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$$

$$\implies \text{Marginal pdf } f_Y(y) = \int_{\underline{\partial}\Omega_1^Y}^{\overline{\partial}\Omega_1^Y} f_{Y,W}(y, w) dw = \int_0^y \frac{(y-w)^{-1/2} w^2 e^{-y/2}}{16\sqrt{2\pi}} dw$$

$$f_Y(y) = \frac{y^{5/2} e^{-y/2}}{8\sqrt{2\pi}} \cdot \int_0^{\pi/2} \left(\frac{1}{2} - \frac{1}{2} \cos 2\theta \right)^2 \sin \theta d\theta$$

$$f_Y(y) \stackrel{TRIG}{=} \frac{y^{5/2} e^{-y/2}}{8\sqrt{2\pi}} \cdot \int_0^{\pi/2} \left[\frac{1}{2} - \frac{1}{2} (2 \cos^2 \theta - 1) \right]^2 \sin \theta d\theta$$

$$f_Y(y) = \frac{y^{5/2} e^{-y/2}}{8\sqrt{2\pi}} \cdot \int_0^{\pi/2} (1 - \cos^2 \theta)^2 \sin \theta d\theta$$

χ_7^2 from χ_6^2 & χ_1^2 pdf's

Theorem

(χ_7^2 Theorem – CHISQ7THM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_6^2 \implies Y := X_1 + X_2 \sim \chi_7^2$ where $f_Y(y) = \frac{y^{5/2} e^{-y/2}}{15\sqrt{2\pi}} = \frac{y^{7/2-1} e^{-y/2}}{2^{7/2} \cdot \Gamma(7/2)} \leftarrow \Gamma_{7/2,2}$ pdf

This distribution of Y is called the **chi-square distribution with seven degrees of freedom**.

PROOF:

$$\implies \text{Joint pdf } f_{Y,W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2} w^2 e^{-y/2}}{16\sqrt{2\pi}}$$

$$\implies \text{Joint Support } \Omega_{Y,W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$$

$$\implies \text{Marginal pdf } f_Y(y) = \int_{\underline{\Omega}_1^Y}^{\overline{\Omega}_1^Y} f_{Y,W}(y, w) dw = \int_0^y \frac{(y-w)^{-1/2} w^2 e^{-y/2}}{16\sqrt{2\pi}} dw$$

$$f_Y(y) = \frac{y^{5/2} e^{-y/2}}{8\sqrt{2\pi}} \cdot \int_0^{\pi/2} (1 - \cos^2 \theta)^2 \sin \theta d\theta = \frac{y^{5/2} e^{-y/2}}{8\sqrt{2\pi}} \cdot \int_0^{\pi/2} (1 - 2\cos^2 \theta + \cos^4 \theta) \sin \theta d\theta$$

$$(CV) u := \cos \theta \iff du = -\sin \theta \implies u(\pi/2) = \cos(\pi/2) = 0, u(0) = \cos(0) = 1$$

$$f_Y(y) \stackrel{CV}{=} \frac{y^{5/2} e^{-y/2}}{8\sqrt{2\pi}} \cdot \int_0^1 (1 - 2u^2 + u^4) du = \frac{y^{5/2} e^{-y/2}}{8\sqrt{2\pi}} \cdot \left[u - \frac{2}{3}u^3 + \frac{1}{5}u^5 \right]_{u=0}^{u=1}$$

$$f_Y(y) \stackrel{FTC}{=} \frac{y^{5/2} e^{-y/2}}{8\sqrt{2\pi}} \cdot \frac{8}{15} = \frac{y^{5/2} e^{-y/2}}{15\sqrt{2\pi}} = \frac{y^{7/2-1} e^{-y/2}}{2^{7/2} \cdot \Gamma(7/2)} \quad \square$$

| CHISQ RV | PDF SIMPLIFIED | PDF GAMMA FORM | KEY INTEGRAL |
|-------------------|---------------------------|-------------------------------------------------------|----------------------------------------------------------|
| $Y \sim \chi_2^2$ | $\frac{1}{2}e^{-y/2}$ | $\frac{y^{2/2-1}e^{-y/2}}{2^{2/2} \cdot \Gamma(2/2)}$ | $\int_0^{\pi/2} \sin^0 \theta d\theta = \frac{\pi}{2}$ |
| $Y \sim \chi_4^2$ | $\frac{1}{4}ye^{-y/2}$ | $\frac{y^{4/2-1}e^{-y/2}}{2^{4/2} \cdot \Gamma(4/2)}$ | $\int_0^{\pi/2} \sin^2 \theta d\theta = \frac{\pi}{4}$ |
| $Y \sim \chi_6^2$ | $\frac{1}{16}y^2e^{-y/2}$ | $\frac{y^{6/2-1}e^{-y/2}}{2^{6/2} \cdot \Gamma(6/2)}$ | $\int_0^{\pi/2} \sin^4 \theta d\theta = \frac{3\pi}{16}$ |
| \vdots | \vdots | \vdots | \vdots |

$$\int_0^{\pi/2} \sin^{2k} \theta d\theta = \frac{\sqrt{\pi}}{2} \cdot \frac{\Gamma(k + \frac{1}{2})}{\Gamma(k + 1)} = \frac{(1)(3)(5) \cdots (2k - 1)}{(2)(4)(6) \cdots (2k)} \cdot \frac{\pi}{2}, \quad k \in \mathbb{Z}_+$$

$$\int_0^{\pi/2} \sin^{2k-2} \theta d\theta = \frac{\sqrt{\pi}}{2} \cdot \frac{\Gamma(k - \frac{1}{2})}{\Gamma(k)} = \frac{(1)(3)(5) \cdots (2k - 3)}{(2)(4)(6) \cdots (2k - 2)} \cdot \frac{\pi}{2}, \quad k \in \mathbb{Z}_+$$

H.B. Dwight, *Tables and Integrals and Other Math. Data*, 4th Ed, 1961. (Ch12, entry 858.44)
 B.O. Peirce, R.M. Foster, *A Short Table of Integrals*, 4th Ed, 1956. (Part II, entry 498)

χ_{2k+1}^2 Pattern

| CHISQ RV | PDF SIMPLIFIED | PDF GAMMA FORM | KEY INTEGRAL |
|-------------------|--------------------------------------------|--------------------------------------------------------|----------------------------------------------------------|
| $Y \sim \chi_3^2$ | $\frac{1}{\sqrt{2\pi}} y^{1/2} e^{-y/2}$ | $\frac{y^{3/2-1} e^{-y/2}}{2^{3/2} \cdot \Gamma(3/2)}$ | $\int_0^{\pi/2} \sin \theta \, d\theta = 1$ |
| $Y \sim \chi_5^2$ | $\frac{1}{3\sqrt{2\pi}} y^{3/2} e^{-y/2}$ | $\frac{y^{5/2-1} e^{-y/2}}{2^{5/2} \cdot \Gamma(5/2)}$ | $\int_0^{\pi/2} \sin^3 \theta \, d\theta = \frac{2}{3}$ |
| $Y \sim \chi_7^2$ | $\frac{1}{15\sqrt{2\pi}} y^{5/2} e^{-y/2}$ | $\frac{y^{7/2-1} e^{-y/2}}{2^{7/2} \cdot \Gamma(7/2)}$ | $\int_0^{\pi/2} \sin^5 \theta \, d\theta = \frac{8}{15}$ |
| \vdots | \vdots | \vdots | \vdots |

$$\int_0^{\pi/2} \sin^{2k+1} \theta \, d\theta = \frac{\sqrt{\pi}}{2} \cdot \frac{\Gamma(k+1)}{\Gamma((k+1) + \frac{1}{2})} = \frac{(2)(4)(6) \cdots (2k)}{(3)(5)(7) \cdots (2k+1)}, \quad k \in \mathbb{Z}_+$$

$$\int_0^{\pi/2} \sin^{2k-1} \theta \, d\theta = \frac{\sqrt{\pi}}{2} \cdot \frac{\Gamma(k)}{\Gamma(k + \frac{1}{2})} = \frac{(2)(4)(6) \cdots (2k-2)}{(3)(5)(7) \cdots (2k-1)}, \quad k \in \mathbb{Z}_+$$

H.B. Dwight, *Tables and Integrals and Other Math. Data*, 4th Ed, 1961. (Ch12, entry 858.44)

B.O. Peirce, R.M. Foster, *A Short Table of Integrals*, 4th Ed, 1956. (Part II, entry 498)

χ_{2k}^2 from χ_{2k-1}^2 & χ_1^2 pdf's

Theorem

(χ^2 Even Theorem – CHISQETHM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_{2k-1}^2 \implies Y := X_1 + X_2 \sim \chi_{2k}^2$ where $f_Y(y) = \frac{y^{(2k)/2-1} e^{-y/2}}{2^{(2k)/2} \cdot \Gamma((2k)/2)}$

This distribution of Y is called the **chi-square distribution with $(2k)$ degrees of freedom**.

PROOF:

pdf's $f_{X_1}(x_1) = \frac{x_1^{-1/2} e^{-x_1/2}}{2^{1/2} \cdot \Gamma(1/2)}$, $f_{X_2}(x_2) = \frac{x_2^{k-3/2} e^{-x_2/2}}{2^{k-1/2} \cdot \Gamma(k-1/2)} \implies$ Supports $\Omega_{X_1} = \Omega_{X_2} = (0, \infty)$

Joint pdf $f_{X_1, X_2}(x_1, x_2) \stackrel{IND}{=} f_{X_1}(x_1) \cdot f_{X_2}(x_2) = \frac{x_1^{-1/2} x_2^{k-3/2} e^{-(x_1+x_2)/2}}{2^{k-1/2} \cdot 2^{1/2} \cdot \Gamma(k-1/2) \cdot \Gamma(1/2)}$

Joint Support $\Omega_{X_1, X_2} \stackrel{IND}{=} \Omega_{X_1} \times \Omega_{X_2} = \{(x_1, x_2) : x_1 \in \Omega_{X_1}, x_2 \in \Omega_{X_2}\} = \{(x_1, x_2) : x_1 > 0, x_2 > 0\}$

Change of random variables (CRV):

$$\begin{aligned} Y := X_1 + X_2 & \iff X_1 = Y - W \\ W := X_2 & \iff X_2 = W \end{aligned} \implies J = \begin{bmatrix} \frac{\partial X_1}{\partial Y} & \frac{\partial X_1}{\partial W} \\ \frac{\partial X_2}{\partial Y} & \frac{\partial X_2}{\partial W} \end{bmatrix} = \begin{bmatrix} 1 & -1 \\ 0 & 1 \end{bmatrix} \implies |J| = 1$$

\implies Joint pdf $f_{Y, W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2} w^{k-3/2} e^{-y/2}}{2^{k-1/2} \cdot 2^{1/2} \cdot \Gamma(k-1/2) \cdot \Gamma(1/2)}$

\implies Joint Support $\Omega_{Y, W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$

χ_{2k}^2 from χ_{2k-1}^2 & χ_1^2 pdf's

Theorem

(χ^2 Even Theorem – CHISQETHM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_{2k-1}^2 \implies Y := X_1 + X_2 \sim \chi_{2k}^2$ where $f_Y(y) = \frac{y^{(2k)/2-1} e^{-y/2}}{2^{(2k)/2} \cdot \Gamma((2k)/2)}$

This distribution of Y is called the **chi-square distribution with $(2k)$ degrees of freedom**.

PROOF:

$$\implies \text{Joint pdf } f_{Y,W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{x_1^{-1/2} x_2^{k-3/2} e^{-(x_1+x_2)/2}}{2^{k-1/2} \cdot 2^{1/2} \cdot \Gamma(k-1/2) \cdot \Gamma(1/2)}$$

$$\implies \text{Joint Support } \Omega_{Y,W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$$

$$\implies \text{Marginal pdf } f_Y(y) = \int_{\underline{\partial}\Omega_1^Y}^{\overline{\partial}\Omega_1^Y} f_{Y,W}(y, w) dw = \int_0^y \frac{(y-w)^{-1/2} w^{k-3/2} e^{-y/2}}{2^{k-1/2} \cdot 2^{1/2} \cdot \Gamma(k-1/2) \cdot \Gamma(1/2)} dw$$

$$f_Y(y) = \int_0^y \frac{(y-w)^{-1/2} w^{k-3/2} e^{-y/2}}{2^{k-1/2} \cdot 2^{1/2} \cdot \Gamma(k-1/2) \cdot \Gamma(1/2)} dw \quad (CV) \quad w := y \sin^2 \theta \Leftrightarrow \begin{cases} y-w = y \cos^2 \theta \\ dw = 2y \sin \theta \cos \theta d\theta \\ w = y \iff \theta = \pi/2 \\ w = 0 \iff \theta = 0 \end{cases}$$

$$f_Y(y) \stackrel{CV}{=} \frac{1}{2^{k-1/2} \cdot 2^{1/2} \cdot \Gamma(k-1/2) \cdot \Gamma(1/2)} \cdot \int_0^{\pi/2} \frac{y^{k-3/2} e^{-y/2} \sin^{2k-3} \theta \cdot (2y \sin \theta \cos \theta d\theta)}{y^{1/2} \cos \theta}$$

χ_{2k}^2 from χ_{2k-1}^2 & χ_1^2 pdf's

Theorem

(χ^2 Even Theorem – CHISQETHM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_{2k-1}^2 \implies Y := X_1 + X_2 \sim \chi_{2k}^2$ where $f_Y(y) = \frac{y^{(2k)/2-1} e^{-y/2}}{2^{(2k)/2} \cdot \Gamma((2k)/2)}$

This distribution of Y is called the **chi-square distribution with $(2k)$ degrees of freedom**.

PROOF:

$$\implies \text{Joint pdf } f_{Y,W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{x_1^{-1/2} x_2^{k-3/2} e^{-(x_1+x_2)/2}}{2^{k-1/2} \cdot 2^{1/2} \cdot \Gamma(k-1/2) \cdot \Gamma(1/2)}$$

$$\implies \text{Joint Support } \Omega_{Y,W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$$

$$\implies \text{Marginal pdf } f_Y(y) = \int_{\underline{\partial}\Omega_Y^y}^{\bar{\partial}\Omega_Y^y} f_{Y,W}(y, w) dw = \int_0^y \frac{(y-w)^{-1/2} w^{k-3/2} e^{-y/2}}{2^{k-1/2} \cdot 2^{1/2} \cdot \Gamma(k-1/2) \cdot \Gamma(1/2)} dw$$

$$f_Y(y) \stackrel{CV}{=} \frac{1}{2^{k-1/2} \cdot 2^{1/2} \cdot \Gamma(k-1/2) \cdot \Gamma(1/2)} \cdot \int_0^{\pi/2} \frac{y^{k-3/2} e^{-y/2} \sin^{2k-3} \theta \cdot (2y \sin \theta \cos \theta d\theta)}{y^{1/2} \cos \theta}$$

$$f_Y(y) = \frac{y^{k-1} e^{-y/2}}{2^{k-3/2} \cdot 2^{1/2} \cdot \Gamma(k-1/2) \cdot \Gamma(1/2)} \cdot \int_0^{\pi/2} \sin^{2k-2} \theta d\theta$$

$$f_Y(y) = \frac{y^{k-1} e^{-y/2}}{2^{k-3/2} \cdot 2^{1/2} \cdot \Gamma(k-1/2) \cdot \Gamma(1/2)} \cdot \left[\frac{\sqrt{\pi}}{2} \cdot \frac{\Gamma(k-1/2)}{\Gamma(k)} \right] = \frac{y^{(2k)/2-1} e^{-y/2}}{2^{(2k)/2} \cdot \Gamma((2k)/2)} \quad \square$$

χ_{2k+1}^2 from χ_{2k}^2 & χ_1^2 pdf's

Theorem

(χ^2 Odd Theorem – CHISQOTHM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_{2k}^2 \implies Y := X_1 + X_2 \sim \chi_{2k+1}^2$ where $f_Y(y) = \frac{y^{(2k+1)/2-1} e^{-y/2}}{2^{(2k+1)/2} \cdot \Gamma((2k+1)/2)}$

This distribution of Y is called the **chi-square distribution with $(2k+1)$ degrees of freedom**.

PROOF:

pdf's $f_{X_1}(x_1) = \frac{x_1^{-1/2} e^{-x_1/2}}{2^{1/2} \cdot \Gamma(1/2)}$, $f_{X_2}(x_2) = \frac{x_2^{k-1} e^{-x_2/2}}{2^k \cdot \Gamma(k)} \implies$ Supports $\Omega_{X_1} = \Omega_{X_2} = (0, \infty)$

Joint pdf $f_{X_1, X_2}(x_1, x_2) \stackrel{IND}{=} f_{X_1}(x_1) \cdot f_{X_2}(x_2) = \frac{x_1^{-1/2} x_2^{k-1} e^{-(x_1+x_2)/2}}{2^k \cdot 2^{1/2} \cdot \Gamma(k) \cdot \Gamma(1/2)}$

Joint Support $\Omega_{X_1, X_2} \stackrel{IND}{=} \Omega_{X_1} \times \Omega_{X_2} = \{(x_1, x_2) : x_1 \in \Omega_{X_1}, x_2 \in \Omega_{X_2}\} = \{(x_1, x_2) : x_1 > 0, x_2 > 0\}$

Change of random variables (CRV):

$$\begin{aligned} Y := X_1 + X_2 \\ W := X_2 \end{aligned} \iff \begin{aligned} X_1 = Y - W \\ X_2 = W \end{aligned} \implies J = \begin{bmatrix} \frac{\partial X_1}{\partial Y} & \frac{\partial X_1}{\partial W} \\ \frac{\partial X_2}{\partial Y} & \frac{\partial X_2}{\partial W} \end{bmatrix} = \begin{bmatrix} 1 & -1 \\ 0 & 1 \end{bmatrix} \implies |J| = 1$$

\implies Joint pdf $f_{Y, W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{(y-w)^{-1/2} w^{k-1} e^{-y/2}}{2^k \cdot 2^{1/2} \cdot \Gamma(k) \cdot \Gamma(1/2)}$

\implies Joint Support $\Omega_{Y, W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$

χ_{2k+1}^2 from χ_{2k}^2 & χ_1^2 pdf's

Theorem

(χ^2 Odd Theorem – CHISQOTHM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_{2k}^2 \implies Y := X_1 + X_2 \sim \chi_{2k+1}^2$ where $f_Y(y) = \frac{y^{(2k+1)/2-1} e^{-y/2}}{2^{(2k+1)/2} \cdot \Gamma((2k+1)/2)}$

This distribution of Y is called the **chi-square distribution with $(2k+1)$ degrees of freedom.**

PROOF:

$$\implies \text{Joint pdf } f_{Y,W}(y,w) \stackrel{CRV}{=} f_{X_1,X_2}(y-w,w) \cdot |J| = \frac{x_1^{-1/2} x_2^{k-1} e^{-(x_1+x_2)/2}}{2^k \cdot 2^{1/2} \cdot \Gamma(k) \cdot \Gamma(1/2)}$$

$$\implies \text{Joint Support } \Omega_{Y,W} = \{(y,w) : 0 < y < \infty, 0 < w < y\}$$

$$\implies \text{Marginal pdf } f_Y(y) = \int_{\underline{\partial}\Omega_1^y}^{\overline{\partial}\Omega_1^y} f_{Y,W}(y,w) dw = \int_0^y \frac{(y-w)^{-1/2} w^{k-1} e^{-y/2}}{2^k \cdot 2^{1/2} \cdot \Gamma(k) \cdot \Gamma(1/2)} dw$$

$$f_Y(y) = \int_0^y \frac{(y-w)^{-1/2} w^{k-1} e^{-y/2}}{2^k \cdot 2^{1/2} \cdot \Gamma(k) \cdot \Gamma(1/2)} dw \quad (CV) \quad w := y \sin^2 \theta \iff \begin{cases} y-w & = y \cos^2 \theta \\ dw & = 2y \sin \theta \cos \theta d\theta \\ w=y & \iff \theta = \pi/2 \\ w=0 & \iff \theta = 0 \end{cases}$$

$$f_Y(y) \stackrel{CV}{=} \frac{1}{2^k \cdot 2^{1/2} \cdot \Gamma(k) \cdot \Gamma(1/2)} \cdot \int_0^{\pi/2} \frac{y^{k-1} \sin^{2k-2} \theta \cdot (2y \sin \theta \cos \theta d\theta)}{y^{1/2} \cos \theta}$$

χ_{2k+1}^2 from χ_{2k}^2 & χ_1^2 pdf's

Theorem

(χ^2 Odd Theorem – CHISQOTHM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_{2k}^2 \implies Y := X_1 + X_2 \sim \chi_{2k+1}^2$ where $f_Y(y) = \frac{y^{(2k+1)/2-1} e^{-y/2}}{2^{(2k+1)/2} \cdot \Gamma((2k+1)/2)}$

This distribution of Y is called the **chi-square distribution with $(2k+1)$ degrees of freedom.**

PROOF:

$$\implies \text{Joint pdf } f_{Y,W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{x_1^{-1/2} x_2^{k-1} e^{-(x_1+x_2)/2}}{2^k \cdot 2^{1/2} \cdot \Gamma(k) \cdot \Gamma(1/2)}$$

$$\implies \text{Joint Support } \Omega_{Y,W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$$

$$\implies \text{Marginal pdf } f_Y(y) = \int_{\underline{\partial}\Omega_Y^y}^{\overline{\partial}\Omega_Y^y} f_{Y,W}(y, w) dw = \int_0^y \frac{(y-w)^{-1/2} w^{k-1} e^{-y/2}}{2^k \cdot 2^{1/2} \cdot \Gamma(k) \cdot \Gamma(1/2)} dw$$

$$f_Y(y) \stackrel{CV}{=} \frac{1}{2^k \cdot 2^{1/2} \cdot \Gamma(k) \cdot \Gamma(1/2)} \cdot \int_0^{\pi/2} \frac{y^{k-1} \sin^{2k-2} \theta \cdot (2y \sin \theta \cos \theta d\theta)}{y^{1/2} \cos \theta}$$

$$f_Y(y) = \frac{y^{k-1/2} e^{-y/2}}{2^{k-1} \cdot 2^{1/2} \cdot \Gamma(k) \cdot \Gamma(1/2)} \cdot \int_0^{\pi/2} \sin^{2k-1} \theta d\theta$$

$$f_Y(y) = \frac{y^{k-1/2} e^{-y/2}}{2^{k-1} \cdot 2^{1/2} \cdot \Gamma(k) \cdot \Gamma(1/2)} \cdot \left[\frac{\sqrt{\pi}}{2} \cdot \frac{\Gamma(k)}{\Gamma\left(k + \frac{1}{2}\right)} \right]$$

Theorem

(χ^2 Odd Theorem – CHISQOTHM)

$X_1, X_2 \stackrel{IND}{\sim} \chi_1^2, \chi_{2k}^2 \implies Y := X_1 + X_2 \sim \chi_{2k+1}^2$ where $f_Y(y) = \frac{y^{(2k+1)/2-1} e^{-y/2}}{2^{(2k+1)/2} \cdot \Gamma((2k+1)/2)}$

This distribution of Y is called the **chi-square distribution with $(2k+1)$ degrees of freedom**.

PROOF:

$$\implies \text{Joint pdf } f_{Y,W}(y, w) \stackrel{CRV}{=} f_{X_1, X_2}(y-w, w) \cdot |J| = \frac{x_1^{-1/2} x_2^{k-1} e^{-(x_1+x_2)/2}}{2^k \cdot 2^{1/2} \cdot \Gamma(k) \cdot \Gamma(1/2)}$$

$$\implies \text{Joint Support } \Omega_{Y,W} = \{(y, w) : 0 < y < \infty, 0 < w < y\}$$

$$\implies \text{Marginal pdf } f_Y(y) = \int_{\underline{\partial}\Omega_1^Y}^{\overline{\partial}\Omega_1^Y} f_{Y,W}(y, w) dw = \int_0^y \frac{(y-w)^{-1/2} w^{k-1} e^{-y/2}}{2^k \cdot 2^{1/2} \cdot \Gamma(k) \cdot \Gamma(1/2)} dw$$

$$f_Y(y) = \frac{y^{k-1/2} e^{-y/2}}{2^{k-1} \cdot 2^{1/2} \cdot \Gamma(k) \cdot \Gamma(1/2)} \cdot \left[\frac{\sqrt{\pi}}{2} \cdot \frac{\Gamma(k)}{\Gamma\left(k + \frac{1}{2}\right)} \right]$$

$$f_Y(y) = \frac{y^{k-1/2} e^{-y/2}}{2^k \cdot 2^{1/2} \cdot \Gamma\left(k + \frac{1}{2}\right)} = \frac{y^{(2k+1)/2-1} e^{-y/2}}{2^{(2k+1)/2} \cdot \Gamma((2k+1)/2)} \quad \square$$

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Fin.